R07- Multiple (Linear) Regression

HCI/PSYCH 522 Iowa State University

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Multiple regression

Recall the simple linear regression model is

$$Y_i \stackrel{ind}{\sim} N(\mu_i, \sigma^2), \quad \mu_i = \beta_0 + \beta_1 X_i$$

The multiple regression model has mean

$$\mu_i = \beta_0 + \beta_1 X_{i,1} + \dots + \beta_p X_{i,p}$$

where for observation i

- Y_i is the dependent variable and
- $X_{i,p}$ is the p^{th} independent variable.

independent variables

There is a lot of flexibility in the mean

$$\mu_i = \beta_0 + \beta_1 X_{i,1} + \dots + \beta_p X_{i,p}$$

as there are many possibilities for the independent variables $X_{i,1}, \ldots, X_{i,p}$:

- Functions (f(X))
- Dummy variables for categorical variables ($X_1 = I()$)
- Higher order terms (X^2)
- Additional independent variables (X_1, X_2)
- Interactions (X_1X_2)
 - Categorical-categorical
 - Continuous-categorical
 - Continuous-continuous

Parameter interpretation

Model:

$$Y_i \stackrel{ind}{\sim} N(\beta_0 + \beta_1 X_{i,1} + \dots + \beta_p X_{i,p}, \sigma^2)$$

The interpretation is

- β_0 is the mean value of the dependent variable Y_i when all independent variables are zero.
- β_p , $p \neq 0$ is the mean increase in the dependent variable for a one-unit increase in the p^{th} independent variable when all other independent variables are held constant.
- R^2 is the proportion of the variability in the dependent variable explained by the model

Galileo experiment



Higher order terms (X^2)

Galileo data (Sleuth3::case1001)



Higher order terms (X^2)

Let

- Y_i be the distance for the i^{th} run of the experiment and
- H_i be the height for the i^{th} run of the experiment.

Simple linear regression assumes

$$Y_i \stackrel{ind}{\sim} N(\beta_0 + \beta_1 H_i \qquad , \sigma^2)$$

The quadratic multiple regression assumes

$$Y_i \stackrel{ind}{\sim} N(\beta_0 + \beta_1 H_i + \beta_2 H_i^2 \qquad , \sigma^2)$$

The cubic multiple regression assumes

$$Y_i \stackrel{ind}{\sim} N(\beta_0 + \beta_1 H_i + \beta_2 H_i^2 + \beta_3 H_i^3, \sigma^2)$$

R code and output

m1 = m2 =	<pre>lm(Distance lm(Distance</pre>	~ Height + I(H	Height ²),	case1001) case1001) eight ³), case1001)	
	ficients(m1)				
	Intercept) 269.712458	0			
coef	ficients(m2)				
	*	0	I(Height^2) -3.436937e-04		
coef	ficients(m3)				
## ##	(Intercept) 1.557755e+02		I(Height^2) -1.244943e-03	. 0	

Higher order terms (X^2)

Galileo experiment (Sleuth3::case1001)



Longnose Dace Abundance

From http://udel.edu/~mcdonald/statmultreg.html:

I extracted some data from the Maryland Biological Stream Survey. ... The [dependent] variable is the number of Longnose Dace ... per 75-meter section of [a] stream. The [independent] variables are ... the maximum depth (in cm) of the 75-meter segment of stream; nitrate concentration (mg/liter)

Consider the model

$$Y_i \stackrel{ind}{\sim} N(\beta_0 + \beta_1 X_{i,1} + \beta_2 X_{i,2}, \sigma^2)$$

where

- Y_i : count of Longnose Dace in stream i
- $X_{i,1}$: maximum depth (in cm) of stream i
- $X_{i,2}$: nitrate concentration (mg/liter) of stream i

Exploratory



R code and output

```
m <- lm(count ~ maxdepth + no3, longnosedace)</pre>
summary(m)
##
## Call:
## lm(formula = count ~ maxdepth + no3, data = longnosedace)
##
## Residuals:
##
      Min
               10 Median
                               30
                                     Max
## -55.060 -27.704 -8.679 11.794 165.310
##
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -17.5550 15.9586 -1.100 0.27544
## maxdepth 0.4811 0.1811 2.656 0.00997 **
## no3
                8,2847 2,9566
                                   2.802 0.00671 **
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 43.39 on 64 degrees of freedom
## Multiple R-squared: 0.1936, Adjusted R-squared: 0.1684
## F-statistic: 7.682 on 2 and 64 DF, p-value: 0.001022
```

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R code and output

ci	<- confint(r	n)	
ci			
##		2.5 %	97.5 %
##	(Intercept)	-49.4361015	14.3260349
##	maxdepth	0.1192458	0.8428725
##	no3	2.3782494	14.1912007

Interpretation

- Intercept (β₀): The mean count of Longnose Dace when maximum depth and nitrate concentration are both zero is -18 (-49, 14).
- Coefficient for maxdepth (β₁): Holding nitrate concentration constant, each cm increase in maximum depth is associated with an additional 0.48 (0.12, 0.84) Longnose Dace counted on average.
- Coefficient for no3 (β₂): Holding maximum depth constant, each mg/L increase in nitrate concentration is associated with an addition 8.3 (2.4, 14.2) Longnose Dace counted on average.
- Coefficient of determination (R²): The model explains 19% of the variability in the count of Longnose Dace.

Interactions

Why an interaction?

Two independent variables are said to interact if the effect that one of them has on the mean dependent variable depends on the value of the other.

For example,

- Crop yield: the effect of tillage method depends on the fertilizer brand (Categorical-categorical)
- Energy expenditure: The effect of mass depends on the species type. (Continuous-categorical)
- Longnose dace count: The effect of nitrate (no3) on longnose dace count depends on the maxdepth. (Continuous-continuous)

Seaweed regeneration (Sleuth3::case1301 subset)



Categorical-categorical

Let category A and type 0 be the reference level. For observation i, let

- Y_i be the dependent variable,
- 1_i be a dummy variable for type 1,
- B_i be a dummy variable for category B, and
- C_i be a dummy variable for category C.

The mean containing only main effects is

$$\mu_i = \beta_0 + \beta_1 1_i + \beta_2 B_i + \beta_3 C_i.$$

The mean with an interaction is

$$\mu_i = \beta_0 + \beta_1 1_i + \beta_2 B_i + \beta_3 C_i + \beta_4 1_i B_i + \beta_5 1_i C_i.$$

Interpretation for the main effects model

The mean containing only main effects is

$$\mu_i = \beta_0 + \beta_1 1_i + \beta_2 B_i + \beta_3 C_i.$$

The means in the main effect model are

Interpretation for the model with an interaction

The mean with an interaction is

$$\mu_i = \beta_0 + \beta_1 1_i + \beta_2 B_i + \beta_3 C_i + \beta_4 1_i B_i + \beta_5 1_i C_i.$$

The means are



This is equivalent to a cell-means model where each combination has its own mean.

R code and output - main effects only

```
##
## Call.
## lm(formula = Cover ~ Block + Treat, data = case1301_subset)
##
## Residuals.
      Min
##
              10 Median
                             30
                                    Max
## -2.3333 -0.6667 0.0000 0.7917 1.8333
##
## Coefficients:
             Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 4.6667
                        0.7683 6.074 0.000298 ***
## BlockB2
          2.1667 0.7683 2.820 0.022491 *
## TreatLf -1.5000 0.9410 -1.594 0.149578
## TreatLfF -3.0000 0.9410 -3.188 0.012838 *
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.331 on 8 degrees of freedom
## Multiple R-squared: 0.6937, Adjusted R-squared: 0.5788
## F-statistic: 6.039 on 3 and 8 DF, p-value: 0.01881
```

Credible intervals

##		2.5 %	97.5 %
##	(Intercept)	2.8949744	6.4383590
##	BlockB2	0.3949744	3.9383590
##	TreatLf	-3.6698711	0.6698711
##	TreatLfF	-5.1698711	-0.8301289

Interpretation

- In block B1 and treatment L, the mean cover is 4.7 (2.9, 6.4).
- The mean increase in cover from block B1 to block B2 is 2.2 (0.4, 3.9) while holding treatment constant.
- The mean increase in cover from treatment L to treatment Lf is -1.5 (-3.7, 0.7) while holding block constant.
- The mean increase in cover from treatment L to treatment LfF is -3 (-5.2, -0.8) while holding block constant.
- The model with block and treatment explains 69% of the variability in cover.

R code and output - with an interaction

```
##
## Call:
## lm(formula = Cover ~ Block * Treat, data = case1301 subset)
##
## Residuals:
##
     Min
             10 Median
                          30
                                Max
## -1.500 -0.625 0.000 0.625
                             1.500
##
## Coefficients:
##
                    Estimate Std. Error t value Pr(>|t|)
##
  (Intercept)
                   4.000e+00 8.898e-01 4.496
                                                0.00412 **
## BlockB2
                   3.500e+00
                             1.258e+00 2.782 0.03193 *
## TreatLf
               -1.976e-15 1.258e+00 0.000 1.00000
           -2.500e+00 1.258e+00
                                         -1.987
## TreatLfF
                                                0.09413
## BlockB2:TreatLf -3.000e+00 1.780e+00
                                         -1.686 0.14280
## BlockB2:TreatLfF -1.000e+00 1.780e+00
                                        -0.562 0.59450
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.258 on 6 degrees of freedom
## Multiple R-squared: 0.7946, Adjusted R-squared: 0.6234
## F-statistic: 4.642 on 5 and 6 DF, p-value: 0.04429
```

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Credible intervals

##		2.5 %	97.5 %
##	(Intercept)	1.8228442	6.1771558
##	BlockB2	0.4210368	6.5789632
##	TreatLf	-3.0789632	3.0789632
##	TreatLfF	-5.5789632	0.5789632
##	BlockB2:TreatLf	-7.3543116	1.3543116
##	BlockB2:TreatLfF	-5.3543116	3.3543116

Interpretation

- In block B1 and treatment L, the mean cover is 4 (1.8, 6.2).
- The mean increase in cover from block B1 to block B2 is 3.5 (0.4, 6.6) in treatment L.
- The mean increase in cover from treatment L to treatment Lf is 0 (-3.1, 3.1) in block B1.
- The mean increase in cover from treatment L to treatment LfF is -2.5 (-5.6, 0.6) in block B1.
- The model that includes block, treatment, and their interaction explains 79% of the variability in cover.

Visualizing the models



In-flight energy expenditure (Sleuth3::case1002)



Continuous-categorical interaction

Continuous-categorical interaction

Let category A be the reference level. For observation i, let

- Y_i be the dependent variable
- $X_{i,1}$ be the continuous independent variable,
- B_i be a dummy variable for category B, and
- C_i be a dummy variable for category C.

The mean containing only main effects is

$$\mu_i = \beta_0 + \beta_1 X_{i,1} + \beta_2 B_i + \beta_3 C_i.$$

The mean with the interaction is

$$\mu_i = \beta_0 + \beta_1 X_{i,1} + \beta_2 B_i + \beta_3 C_i + \beta_4 X_{i,1} B_i + \beta_5 X_{i,1} C_i.$$

Interpretation for the main effect model

The mean containing only main effects is

$$\mu_i = \beta_0 + \beta_1 X_{i,1} + \beta_2 B_i + \beta_3 C_i.$$

For each category, the line is

Category	Line	(μ)	
A	β_0	+	$\beta_1 X$
B	$ \begin{array}{c} (\beta_0 + \beta_2) \\ (\beta_0 + \beta_3) \end{array} $	+	$\beta_1 X$
C	$(\beta_0 + \beta_3)$	+	$\beta_1 X$

Each category has a different intercept, but a common slope.

Interpretation for the model with an interaction

The model with an interaction is

$$\mu_i = \beta_0 + \beta_1 X_{i,1} + \beta_2 B_i + \beta_3 C_i + \beta_4 X_{i,1} B_i + \beta_5 X_{i,1} C_i$$

For each category, the line is

Category	Line (μ)		
A	β_0	$+ \beta_1 \qquad X$	
B	$(\beta_0 + \beta_2)$	$+(\beta_1+\beta_4)X$	
C	$(\beta_0 + \beta_3)$	$+(\beta_1+\beta_5)X$	

Each category has its own intercept and its own slope.

R code and output - main effects only

summary(mM <- lm(log(Energy) ~ log(Mass) + Type, case1002))</pre>

```
##
## Call:
## lm(formula = log(Energy) ~ log(Mass) + Type, data = case1002)
##
## Residuals:
##
       Min
                 10 Median
                                  30
                                          Max
## -0.23224 -0.12199 -0.03637 0.12574 0.34457
##
##
  Coefficients:
##
                            Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                            -1.49770 0.14987 -9.993 2.77e-08 ***
## log(Mass)
                            0.81496 0.04454 18.297 3.76e-12 ***
## Typenon-echolocating bats -0.07866 0.20268 -0.388 0.703
## Typenon-echolocating birds 0.02360 0.15760 0.150 0.883
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.186 on 16 degrees of freedom
## Multiple R-squared: 0.9815, Adjusted R-squared: 0.9781
## F-statistic: 283.6 on 3 and 16 DF, p-value: 4.464e-14
```

```
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```

Credible intervals

##	2.5 %	97.5 %
## (Intercept)	-1.8154046	-1.1799884
## log(Mass)	0.7205339	0.9093811
## Typenon-echolocating bats	-0.5083245	0.3509972
## Typenon-echolocating birds	-0.3104999	0.3576964

Interpretation

- For echo-locating bats that weigh 1 gram (log(Mass) is 0), the mean energy expenditure is -1.5 (-1.8, -1.2) watts.
- The mean increase in the logarithm of energy expenditure per unit increase in the logarithm of mass is 0.8 (0.7, 0.9) while holding type of bird or bat constant.
- The mean increase in the logarithm of energy expenditure from echolocating bats to non-echolocating bats is -0.1 (-0.5, 0.4) while mass constant.
- The mean increase in the logarithm of energy expenditure from echolocating bats to non-echolocating birds is 0 (-0.3, 0.4) while mass constant.
- The model the includes main effects for the logarithm of mass and the type of bird or bat explains 98% of the variability in the logarithm of energy expenditure.

R code and output - with an interaction

summary(mI <- lm(log(Energy) ~ log(Mass) * Type, case1002))</pre>

```
##
## Call:
## lm(formula = log(Energy) ~ log(Mass) * Type, data = case1002)
##
## Residuals:
##
       Min
                 10
                      Median
                                  30
                                          Max
## -0.25152 -0.12643 -0.00954 0.08124 0.32840
##
##
  Coefficients:
##
                                      Estimate Std. Error t value Pr(>|t|)
##
  (Intercept)
                                      -1.47052 0.24767 -5.937 3.63e-05 ***
## log(Mass)
                                                  0.08668 9.283 2.33e-07 ***
                                       0.80466
## Typenon-echolocating bats
                                   1,26807
                                                 1,28542 0,987
                                                                    0.341
## Typenon-echolocating birds
                                      -0.11032
                                                  0.38474 -0.287 0.779
## log(Mass):Typenon-echolocating bats -0.21487
                                                  0.22362 -0.961
                                                                 0.353
## log(Mass):Typenon-echolocating birds 0.03071
                                                  0.10283
                                                           0.299
                                                                    0.770
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.1899 on 14 degrees of freedom
```

Credible intervals

##	2.5 %	97.5 %
## (Intercept)	-2.0017153	-0.9393152
## log(Mass)	0.6187372	0.9905769
## Typenon-echolocating bats	-1.4888841	4.0250195
## Typenon-echolocating birds	-0.9355124	0.7148674
<pre>## log(Mass):Typenon-echolocating bats</pre>	-0.6944979	0.2647479
<pre>## log(Mass):Typenon-echolocating birds</pre>	-0.1898417	0.2512682

Interpretation

- For echo-locating bats that weigh 1 gram (log(Mass) is 0), the mean logarithm of energy expenditure is -1.5 (-2, -0.9) watts.
- The mean increase in the logarithm of energy expenditure per unit increase in the logarithm of mass is 0.8 (0.6, 1) for echolocating bats.
- The mean increase in the logarithm of energy expenditure from echolocating bats to non-echolocating bats is 1.3 (-1.5, 4) when mass is 1 (log(Mass is 0).
- The mean increase in the logarithm of energy expenditure from echolocating bats to non-echolocating birds is -0.1 (-0.9, 0.7) when mass is 1 (log(Mass is 0).
- The model that includes the logarithm of mass, the type of bird and bat, and their interaction explains 98% of the variability in the logarithm of energy.
Visualizing the models



Type — echolocating bats ---- non-echolocating bats --- non-echolocating birds

Continuous-continuous interaction

For observation i, let

- Y_i be the dependent variable
- $X_{i,1}$ be the first continuous independent variable and
- $X_{i,2}$ be the second continuous independent variable.

The mean containing only main effects is

$$\mu_i = \beta_0 + \beta_1 X_{i,1} + \beta_2 X_{i,2}.$$

The mean with the interaction is

$$\mu_i = \beta_0 + \beta_1 X_{i,1} + \beta_2 X_{i,2} + \beta_3 X_{i,1} X_{i,2}.$$

Intepretation - main effects only

Let $X_{i,1} = x_1$ and $X_{i,2} = x_2$, then we can rewrite the line (μ) as

 $\mu = (\beta_0 + \beta_2 x_2) + \beta_1 x_1$

which indicates that the intercept of the line for x_1 depends on the value of x_2 .

Similarly,

$$\mu = (\beta_0 + \beta_1 x_1) + \beta_2 x_2$$

which indicates that the intercept of the line for x_2 depends on the value of x_1 .

Intepretation - with an interaction

Let $X_{i,1} = x_1$ and $X_{i,2} = x_2$, then we can rewrite the mean (μ) as

$$\mu = (\beta_0 + \beta_2 x_2) + (\beta_1 + \beta_3 x_2) x_1$$

which indicates that both the intercept and slope for x_1 depend on the value of x_2 .

Similarly,

$$\mu = (\beta_0 + \beta_1 x_1) + (\beta_2 + \beta_3 x_1) x_2$$

which indicates that both the intercept and slope for x_2 depend on the value of x_1 .

R code and output - main effects only

```
##
## Call:
## lm(formula = count ~ no3 + maxdepth, data = longnosedace)
##
## Residuals:
##
      Min
           10 Median
                             30
                                    Max
## -55,060 -27,704 -8,679 11,794 165,310
##
## Coefficients:
             Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) -17.5550 15.9586 -1.100 0.27544
## no3
       8.2847 2.9566 2.802 0.00671 **
## maxdepth 0.4811 0.1811 2.656 0.00997 **
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 43.39 on 64 degrees of freedom
## Multiple R-squared: 0.1936, Adjusted R-squared: 0.1684
## F-statistic: 7.682 on 2 and 64 DF, p-value: 0.001022
```

Credible intervals

##		2.5 %	97.5 %
##	(Intercept)	-49.4361015	14.3260349
##	no3	2.3782494	14.1912007
##	maxdepth	0.1192458	0.8428725

Interpretation

- When nitrate and maximum depth are both zero, the mean longnosedace count is -17.6 (-49.4, 14.3).
- When nitrate increases by 1 mg/L, the mean longnosedace count increases by 8.3 (2.4, 14.2) while holding maximum depth constant.
- When maximum depth increases by 1 cm, the mean longnosedace count increases by 0.5 (0.1, 0.8) while holding nitrate constant.
- The main effects model explains 19% of the variability in longnose dace count.

R code and output - with an interaction

```
##
## Call.
## lm(formula = count ~ no3 * maxdepth, data = longnosedace)
##
## Residuals.
##
      Min
              10 Median 30
                                     Max
## -65.111 -21.399 -9.562
                           5.953 151.071
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept) 13.321043 23.455710 0.568
                                           0.5721
## no3
          -4.646272 7.856932 -0.591
                                           0.5564
## maxdepth -0.009338 0.329180 -0.028
                                           0.9775
## no3:maxdepth 0.201219 0.113576 1.772
                                            0.0813 .
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 42.68 on 63 degrees of freedom
## Multiple R-squared: 0.2319, Adjusted R-squared: 0.1953
## F-statistic: 6.339 on 3 and 63 DF, p-value: 0.0007966
```

Credible intervals

##		2.5 %	97.5 %
##	(Intercept)	-33.55145355	60.1935389
##	no3	-20.34709813	11.0545539
##	maxdepth	-0.66715251	0.6484768
##	no3:maxdepth	-0.02574574	0.4281832

Interpretation

- When nitrate and maximum depth are both zero, the mean longnosedace count is 13.3 (-33.6, 60.2).
- When nitrate increases by 1 mg/L, the mean longnosedace count increases by -4.6 (-20.3, 11.1) when maximum depth is zero.
- When maximum depth increases by 1 cm, the mean longnosedace count increases by 0 (-0.7, 0.6) when nitrate is zero.
- The model with maximum depth, nitrate, and their interaction explains 23% of the variability in longnose dace count.

Visualizing the model



When to include interaction terms

From The Statistical Sleuth (3rd ed) page 250:

- when a question of interest pertains to an interaction
- when good reason exists to suspect an interaction or
- when interactions are proposed as a more general model for the purpose of examining the goodness of fit of a model without interaction.

Multiple regression independent variables

The possibilities for independent variables are

- Higher order terms (X^2)
- Additional independent variables $(X_1 \text{ and } X_2)$
- Dummy variables for categorical variables $(X_1 = I())$
- Interactions (X_1X_2)
 - Categorical-categorical
 - Continuous-categorical
 - Continuous-continuous

We can also combine these independent variables, e.g.

- including higher order terms for continuous variables along with dummy variables for categorical variables and
- including higher order interactions $(X_1X_2X_3)$.